

# Characterization through Hazard Rate of heavy tailed distributions and some Convolution Closure Properties

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## INTRODUCTION

We will assume for the whole paper that the distribution function  $F$  is supported on  $[0, \infty)$  and that  $F$  has positive Lebesgue density function  $f$ .

We use the Matuszewska indices and its properties to show asymptotics inequalities for the hazard rates. We discuss about the relation membership in dominatedly or subversively varying tail distribution and a hazard rate condition. Convolution closure is establish for the class of distributions with subexponential and subversiverly varying tails.

## The class $\mathcal{L}$

A distribution function  $F$  belongs to the class  $\mathcal{L}$  if

$$\lim_{x \rightarrow \infty} \frac{\overline{F}(x-y)}{\overline{F}(x)} = 1$$

for all constants  $y \in \mathbb{R}$ .

This distribution function  $F$  is said to have a long tail.

## The class $\mathcal{D}$

A distribution function  $F$  belongs to the class  $\mathcal{D}$  if

$$\limsup_{x \rightarrow \infty} \frac{\bar{F}(ux)}{\bar{F}(x)} < \infty \quad (1)$$

for any  $0 < u < 1$  (or equivalently for  $u = 1/2$ ). Such a distribution function  $F$  is said to have a dominatedly varying tail.

An equivalent way to right (1) is

$$\bar{F}_*(u) := \liminf_{x \rightarrow \infty} \frac{\bar{F}(ux)}{\bar{F}(x)} > 0 \quad (2)$$

for any  $u > 1$ .

## The class $\mathcal{E}$

A distribution function  $F$  belongs to the class  $\mathcal{E}$ , if for some  $u > 1$

$$\bar{F}^*(u) := \limsup_{x \rightarrow \infty} \frac{\bar{F}(ux)}{\bar{F}(x)} < 1.$$

Such a distribution function  $F$  is said to have subversively varying tail.

## The class $\mathcal{S}$

A distribution function  $F$  belongs to the class  $\mathcal{S}$  if

$$\lim_{x \rightarrow \infty} \frac{\overline{F^{*n}}(x)}{\overline{F}(x)} = n$$

for any  $n \geq 2$  (or equivalently for  $n = 2$ ), where  $F^{*n}$  denotes the  $n$ th convolution of  $F$ . Such a distribution function  $F$  is said to have a subexponential tail.

## The class $\mathcal{A}$

A distribution function  $F$  belongs to the class  $\mathcal{A}$  if

$$F \in \mathcal{S} \text{ and } \bar{F}^*(u) < 1,$$

for some  $u > 1$ . In others words  $\mathcal{A} = \mathcal{E} \cap \mathcal{S}$ .

It is well known the following inclusions

$$\mathcal{D} \cap \mathcal{A} \subset \mathcal{D} \cap \mathcal{L} \subset \mathcal{S} \subset \mathcal{L}.$$

$$\mathcal{D} \cap \mathcal{A} \subset \mathcal{D} \cap \mathcal{E}$$



## Definition of Matuszewska indices

### Upper Matuszewska index

Let  $g$  be positive. Then

The upper Matuszewska index  $\alpha(g)$  is the infimum of those  $\alpha$  for which there exists a constant  $C = C(\alpha) > 0$  such that for each  $\Lambda > 1$

$$\frac{g(\lambda x)}{g(x)} \leq C(1 + o(1))\lambda^\alpha,$$

as  $x \rightarrow \infty$  and uniformly in  $\lambda \in [1, \Lambda]$ .

## Lower Matuszewska index

Let  $g$  be positive. Then

The lower Matuszewska index  $\beta(g)$  is the supremum of those  $\beta$  for which there exists a constant  $D = D(\beta) > 0$  such that for each  $\Lambda > 1$

$$\frac{g(\lambda x)}{g(x)} \geq D(1 + o(1))\lambda^\beta,$$

as  $x \rightarrow \infty$ , and uniformly in  $\lambda \in [1, \Lambda]$ .

Let consider the following notation. The positive function  $g$  has

- ① bounded increase, ( $g \in BI$ ), if  $\alpha(g) < \infty$
- ② bounded decrease, ( $g \in BD$ ), if  $\beta(g) > -\infty$
- ③ positive increase, ( $g \in PI$ ), if  $\beta(g) > 0$
- ④ positive decrease, ( $g \in PD$ ), if  $\alpha(g) < 0$

An equation that holds between the Matusweska indices is

$$\beta(g) = -\alpha\left(\frac{1}{g}\right), \quad \text{for } g \text{ positive} \quad (3)$$

For more details of the Matuszewska indices, see Chapter 2.1 of Bingham *et al.* (1987).

## Potter Inequalities

Let  $g(\cdot)$  be positive.

If  $g \in \text{BI}$  then for every  $\alpha > \alpha(g)$  there exist positive constants  $C, x_0$  such that

$$\frac{g(y)}{g(x)} \leq C \left(\frac{y}{x}\right)^\alpha, \quad (y \geq x \geq x_0) \quad (4)$$

If  $g \in \text{BD}$  then for every  $\beta < \beta(g)$  there exist positive constants  $C', x_0$  such that

$$\frac{g(y)}{g(x)} \geq C' \left(\frac{y}{x}\right)^\beta, \quad (y \geq x \geq x_0) \quad (5)$$

## Matuszewska indices for the distributions tails

From D. B. H. Cline and G. Samorodnisky(1994)

The upper Matuszewska index  $\gamma_{\bar{F}}$  for a distribution function  $F$ , was introduced as follows.

$$\gamma_{\bar{F}} := \inf \left\{ -\frac{\log \bar{F}_*(u)}{\log u} : u > 1 \right\} = - \lim_{u \rightarrow \infty} \frac{\log \bar{F}_*(u)}{\log u}. \quad (6)$$

The lower Matuszewska index  $\delta_{\bar{F}}$  for a distribution function  $F$ , was introduced as follows.

$$\delta_{\bar{F}} := \sup \left\{ -\frac{\log \bar{F}^*(u)}{\log u} : u > 1 \right\} = - \lim_{u \rightarrow \infty} \frac{\log \bar{F}^*(u)}{\log u}. \quad (7)$$

## Inequalities for the distributions functions

If the Upper Matuszeska index  $\gamma_{\bar{F}} < \infty$  then there exist constants  $C, x_0$  such that

$$\frac{\bar{F}(x)}{\bar{F}(y)} \leq C \left(\frac{x}{y}\right)^{-\gamma} \quad (8)$$

for every  $x \geq y \geq x_0$  and  $\gamma_{\bar{F}} < \gamma < \infty$ .

If the Lower Matuszeska index  $\delta_{\bar{F}} > 0$  is finite then there exist constants  $C', x_0$  such that

$$\frac{\bar{F}(x)}{\bar{F}(y)} \geq C' \left(\frac{x}{y}\right)^{-\delta} \quad (9)$$

for every  $x \geq y \geq x_0$  and  $0 < \delta < \delta_{\bar{F}}$ .

## Matuszewska indices for the density functions

We introduced the upper Matuszewska index  $\gamma_f$  for the density function, as follows

$$\gamma_f := \inf \left\{ -\frac{\log f_*(u)}{\log u} : u > 1 \right\} = -\lim_{u \rightarrow \infty} \frac{\log f_*(u)}{\log u}. \quad (10)$$

where  $f_*(u) = \liminf_{x \rightarrow \infty} \frac{f(ux)}{f(x)}$ .

We introduced the lower Matuszewska index  $\delta_f$  for the density function, as follows

$$\delta_f := \sup \left\{ -\frac{\log f^*(u)}{\log u} : u > 1 \right\} = -\lim_{u \rightarrow \infty} \frac{\log f^*(u)}{\log u}. \quad (11)$$

where  $f^*(u) = \limsup_{x \rightarrow \infty} \frac{f(ux)}{f(x)}$ .

## Definition

We will say that the density function has

- 1 bounded decrease, ( $f \in BD$ ), if  $\gamma_f = \alpha(f^{-1}) < \infty$
- 2 positive decrease, ( $f \in PD$ ), if  $\delta_f = \beta(f^{-1}) > 0$



In analogy to inequalities (8) and (9) we introduce inequalities for density function

### Corollary

If  $f \in BD$  then there exist constants  $C, x_0$  such that

$$\frac{f(y)}{f(x)} \geq C' \left(\frac{y}{x}\right)^{-\gamma} \quad (12)$$

for every  $y \geq x \geq x_0$ .

If  $f \in PD$  then there exist constants  $C', x_0$  such that

$$\frac{f(y)}{f(x)} \leq C \left(\frac{y}{x}\right)^{-\delta} \quad (13)$$

for every  $y \geq x \geq x_0$ .

## Subversively Varying Tail

The Subversively Class is a large class that extends out of the class of Heavy Tail. We can see the following Example

### Example

The Exponential distribution function is  $\bar{F}(x) = \exp\{-\lambda x\}$ . As we can see

$$\lim_{x \rightarrow \infty} \frac{\bar{F}(ux)}{\bar{F}(x)} = 0 < 1$$

for all  $u > 1$ .

## Lemma (Konstantinides et.al 2002)

Let  $F$  be a d.f. with a density function  $f$  which is eventually non-increasing. Then the following statements are equivalent:

- 1  $\bar{F}^*(u) < 1$  holds for some  $u > 1$ ;
- 2  $\bar{F}^*(u) < 1$  holds for any  $u > 1$ ;
- 3 the hazard rate function of  $F$ ,  $h(x) = \frac{f(x)}{\bar{F}(x)}$  satisfies

$$\liminf_{x \rightarrow \infty} xh(x) > 0.$$

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We can obtain that condition (14) is the Potter inequality (4) for  $C \leq 1$  and  $\alpha(g) < 0$ . Obviously

$$\frac{g(y)}{g(x)} \leq C \left(\frac{y}{x}\right)^\alpha \leq C, \quad (x \geq y \geq x_0). \quad (15)$$

for all  $\alpha(g) < \alpha < 0$ . But to obtain, if the Potter inequality holds we only need to calculate the Upper Matuszewska index



The main idea is to replace the condition of eventually non-increasing(or decreasing) with the Potter Inequalities. When we have density functions we can obtain an analogous inequality of the previous one, if the  $\delta_f > 0$ .

## Theorem

Let  $F$  be an absolute continuous distribution function supported on  $[0, \infty)$  with a density function  $f(x)$  with  $\delta_f > 0$  ( $f \in PD$ ). Then

$$F \in \mathcal{E} \text{ if and only if } \liminf_{x \rightarrow \infty} xh(x) > 0. \quad (16)$$

A sufficient condition for  $F \in \mathcal{E}$  is given from the following theorem

### Theorem

*If  $\liminf_{x \rightarrow \infty} xh(x) > 0$  then  $F \in \mathcal{E}$*

A similar theorem is found in Klüppelberg C.(1988) for the class  $\mathcal{D} \cap \mathcal{L}$ .

## Lemma

If  $f \in PD$ ,  $\delta_f > 1$  for any  $\delta \in (1, \delta_f)$  then :

$$xh(x) = x \frac{f(x)}{\bar{F}(x)} \geq \frac{(\delta - 1)}{C} > 0$$

for all  $x \geq x_0$  and  $C > 0$ .

From Lemma and previous theorem we obtain

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for all  $x \geq x_0$  and  $C > 0$ .

From Lemma and previous theorem we obtain

## Corollary

If  $f \in PD$  with  $\delta_f > 1$  then  $F \in \mathcal{E}$ .

## Theorem

If  $F_1, F_2 \in \mathcal{E}$  and the following statements hold

1  $f_1 \in PD$

2  $x^{-\delta^*} = O\left(\frac{1}{\overline{F_1(x)}}\right)$

where  $\delta^* = \min(\delta_{\overline{F_1}}, \delta_{\overline{F_2}})$ , then  $F_1 * F_2 \in \mathcal{E}$ .

## Subexponential Class

A sufficient condition for  $F \in \mathcal{D} \cap \mathcal{L}$  is given by the following theorem.

### Theorem (Klüppelberg C. 1988)

. If  $\limsup_{x \rightarrow \infty} xh(x) < \infty$  then  $F \in \mathcal{D} \cap \mathcal{L}$

### Corollary (Klüppelberg C. 1988)

Let  $F$  have an eventually decreasing density  $f$ . Then the following statements are equivalent:

- 1  $F \in \mathcal{D}$
- 2  $F \in \mathcal{D} \cap \mathcal{L}$
- 3  $\limsup_{x \rightarrow \infty} xh(x) < \infty$

## Theorem

Let  $F$  be an absolute continuous distribution function supported on  $[0, \infty)$  with a density function  $f(x)$  with  $\delta_f > 0$  ( $f \in PD$ ). Then

- 1  $F \in \mathcal{D}$
- 2  $F \in \mathcal{D} \cap \mathcal{L}$
- 3  $\limsup_{x \rightarrow \infty} xh(x) < \infty$



## Lemma

If  $f \in BD$  then there is positive  $x_0$ , such that for all  $x \geq x_0$  and all  $\lambda > 1$ :

$$xh(x) = x \frac{f(x)}{\bar{F}(x)} \leq \frac{-\gamma + 1}{C'(\lambda^{-\gamma+1} - 1)}$$

Furthermore if  $f \in BD$  then  $F \in \mathcal{D} \cap \mathcal{L}$ .

## The class $\mathcal{A} \cap \mathcal{D}$

### Theorem

Let  $F$  be an absolute continuous distribution function with a density function  $f(x)$  with  $\delta_f > 0$  then  $F \in \mathcal{A} \cap \mathcal{D}$  if and only if one of the following statements holds

- 1  $0 < \liminf_{x \rightarrow \infty} xh(x) \leq \limsup_{x \rightarrow \infty} xh(x) < \infty$
- 2  $0 < \bar{F}_*(u) \leq \bar{F}^*(u) < 1$

## Pitman (A Characterization Theorem for $\mathcal{S}$ )

### Theorem (Pitman(1980))

*Suppose  $F$  is absolutely continuous with density function  $f$  and hazard rate  $h(x)$  eventually decreasing to 0. Then  $F \in \mathcal{S}$  if and only if*

$$\lim_{x \rightarrow \infty} \int_0^x \exp \{yh(x)\} f(y) dy = 1 \quad (17)$$





## Theorem





Suppose  $F$  is absolutely continuous with density function  $f$  and hazard rate  $h(x)$  with  $\alpha(h) < 0$ . Then  $F \in \mathcal{S}$  if and only if





$$\lim_{x \rightarrow \infty} \int_0^x \exp \{kyh(x)\} f(y) dy = 1 \quad (18)$$

for all  $k > 0$ .

## BIBLIOGRAPHY

-  BINGHAM, N.H., GOLDIE, C.M., TEUGELS, J.L. 1987 *Regular Variation*. Cambridge University Press, New York.
-  EMBRECHTS, P., GOLDIE, C.M. (1980) On closure and factorization properties of subexponential and related distributions. *J Austral.Math.Soc.* **29**, no.1, 243–256.
-  CAI JUN (2004) On max-sum equivalence and convolution closure of heavy-tailed distributions and their applications. *J Appl. Prob.* **41**, 117-130.
-  D. B. H. Cline and G. Samorodnisky, Subexponentiality of the product of independent random variables, *Stochastic Processes. Appl.* **49** (1994), no. 1, 75–98.

-  EMBRECHTS, P. KLÜPPELBERG, C. AND MIKOSCH, T. (1997) Modeling Extremal Events for Insurance and Finance. *Springer-Verlag*, Berlin, 1997
-  KLÜPPELBERG C. (1988) Subexponential Distributions and Integrated Tails *J Appl. Prob.* **25**, no.1, 132–141.
-  KONSTANTINIDES D.G. (2008) A Class of Heavy Tailed Distributions *Journal of Numerical and Applied Mathematics* **96**, No.1, 127-138.
-  KONSTANTINIDES D.G., TANG, Q.H, TSITSIASHVILI, G.SH (2002) Estimates for the Ruin Probability in the Classical Risk Model with Constant Interest Force in the Presence of Heavy Tails. *Insurance: Mathematics & Economics* **31**, no.3, 447–460.

-  PITMAN E. J. G. (1980) Subexponential Distribution Function  
*J.Austral.Math.Soc.* **29** ,337–347.
-  PRATT W. JOHN (1960) On Interchanging Limits and integrals  
*Ann. Math. Statist.***31**, 74–77
-  Tang, Q. and Tsitsiashvili,G.(2003a). Randomly weighted sums of subexponential random variables with applications to ruin theory,  
*Extremes* **6**,171–188.
-  TANG QIHE (2006) The Subexponentiality of Products Revisited.  
*Extremes* **9**, 231–241.

**Thank you for your attention**

