

Asymptotics for ruin probabilities in a discrete-time risk model with dependent financial and insurance risks

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1. Discrete-time insurance risk model.

Let $\{X_i, i \geq 1\}$ and $\{Y_i, i \geq 1\}$ be insurance and financial risks and the aggregate net losses

$$S_n = \sum_{i=1}^n X_i \prod_{j=1}^i Y_j, \quad (1)$$

for each positive integer n . The finite and infinite time ruin probabilities are

$$\psi(x, n) = \mathbf{P} \left(\max_{1 \leq k \leq n} S_k > x \right), \quad (2)$$

$$\psi(x) = \lim_{n \rightarrow \infty} \psi(x, n) = \mathbf{P} \left(\sup_{n \geq 1} S_n > x \right), \quad (3)$$

where $x \geq 0$ is interpreted as the initial capital.

If we denote the product $\theta_i = \prod_{j=1}^i Y_j$ in (1), the ruin probabilities $\psi(x, n)$ and $\psi(x)$ represent the tail probabilities of the maximum of randomly weighted sums. In case of independence between $\{X_i, i \geq 1\}$ and $\{\theta_i, i \geq 1\}$, under the presence of heavy-tailed insurance risks, was recently established the asymptotic formula

$$\psi(x, n) \sim \sum_{i=1}^n \mathbf{P}(X_i \theta_i > x), \quad x \rightarrow \infty, \quad (4)$$

holds for each fixed n , or

$$\psi(x) \sim \sum_{i=1}^{\infty} \mathbf{P}(X_i \theta_i > x), \quad x \rightarrow \infty. \quad (5)$$

1. Dependence structure.

The survival copula $\overline{C}(u, v)$ is defined by the formula

$$\overline{C}(u, v) = u + v - 1 + C(1 - u, 1 - v), \quad (u, v) \in [0, 1]^2,$$

with respect to the given copula $C(u, v)$. Clearly, the survival copula with respect to $C(u, v)$ can be represented as

$$\overline{C}(\overline{F}(x), \overline{G}(y)) = \mathbf{P}(X > x, Y > y).$$

Assume that the copula function $C(u, v)$ is absolutely continuous. Denote by $C_1(u, v) = \frac{\partial}{\partial u} C(u, v)$, $C_2(u, v) = \frac{\partial}{\partial v} C(u, v)$, $C_{12}(u, v) = \frac{\partial^2}{\partial u \partial v} C(u, v)$. Then

$$\overline{C}_2(u, v) := \frac{\partial}{\partial v} \overline{C}(u, v) = 1 - C_2(1 - u, 1 - v),$$

and

$$\overline{C}_{12}(u, v) := \frac{\partial^2}{\partial u \partial v} \overline{C}(u, v) = C_{12}(1 - u, 1 - v).$$

Assumption A₁. (Albrecher et al. 06) *There exists a positive constant M such that*

$$\limsup_{v \uparrow 1} \limsup_{u \uparrow 1} C_{12}(u, v) = \limsup_{v \uparrow 1} \limsup_{u \uparrow 1} \overline{C}_{12}(1 - u, 1 - v) < M.$$

Assumption A₂. (Asimit and Badescu 10) *The relation*

$$\overline{C}_2(u, v) \sim u \overline{C}_{12}(0+, v), \quad u \downarrow 0,$$

holds uniformly on (0, 1).

Clearly, Assumption A₂ is equivalent to

$$1 - C_2(u, v) \sim (1 - u) C_{12}(1-, v), \quad u \uparrow 1,$$

holds uniformly on (0, 1). Thus, if the copula $C(u, v)$ of the random vector (X, Y) satisfies Assumptions A₁ and A₂, then the copula of (X^+, Y) , denoted by $C^+(u, v)$, satisfies these two assumptions as well.

Assumption A_3 . *The relation*

$$C_2(u, v) = 1 - \overline{C}_2(1 - u, 1 - v) \rightarrow 0, \quad u \downarrow 0,$$

holds uniformly on $[0, 1]$.

We remark that Assumption A_3 is equivalent to the fact that

$$\mathbf{P}(X \leq x \mid Y = y) = C_2[F(x), G(y)] \rightarrow 0, \quad x \rightarrow -\infty, \quad (6)$$

holds uniformly on \mathbb{R} .

We remind the classes of distributions:

1.

$$\mathcal{L} = \left\{ F \mid \lim_{x \rightarrow \infty} \frac{\overline{F}(x-y)}{\overline{F}(x)} = 1, \forall y \in \mathbb{R} \right\},$$

2.

$$\mathcal{D} = \left\{ F \mid \limsup_{x \rightarrow \infty} \frac{\overline{F}(xu)}{\overline{F}(x)} < \infty, \forall u \in (0, 1) \right\}.$$

3.

$$\mathcal{C} = \left\{ F \mid \lim_{u \uparrow 1} \limsup_{x \rightarrow \infty} \frac{\overline{F}(xu)}{\overline{F}(x)} = 1, \right\}.$$

A distribution F on \mathbb{R} belongs to the class $R_{-\alpha}$, if $\lim \overline{F}(x y) / \overline{F}(x) = y^{-\alpha}$ for some $\alpha \geq 0$ and all $y > 0$. It is well known that the following inclusion relationships hold:

$$R_{-\alpha} \subset \mathcal{C} \subset \mathcal{L} \cap \mathcal{D} \subset \mathcal{L}$$

Furthermore, for a distribution F on \mathbb{R} , denote its upper and lower Matuszewska indices, respectively, by

$$J_F^+ = - \lim_{y \rightarrow \infty} \frac{\log \overline{F}_*(y)}{\log y} \quad \text{with} \quad \overline{F}_*(y) := \liminf \frac{\overline{F}(x y)}{\overline{F}(x)} \quad \text{for } y > 1,$$

$$J_F^- = - \lim_{y \rightarrow \infty} \frac{\log \overline{F}^*(y)}{\log y} \quad \text{with} \quad \overline{F}^*(y) := \limsup \frac{\overline{F}(x y)}{\overline{F}(x)} \quad \text{for } y > 1.$$

Clearly, $F \in \mathcal{D}$ is equivalent to $J_F^+ < \infty$. For each $i \geq 1$, denote by H_i the distribution of $X_i \prod_{j=1}^i Y_j$, $j \geq 1$ and let $H_1 = H$.

3. Asymptotic results.

Theorem 1. *In the discrete-time risk model, assume that $\{(X_i, Y_i), i \geq 1\}$ are i.i.d. random vectors with generic random vector (X, Y) satisfying Assumptions A_1 – A_3 . If $F \in \mathcal{C}$ and $\mathbf{E}Y^p < \infty$ for some $p > J_F^+$, then, for each fixed $n \geq 1$, it holds that*

$$\psi(x, n) \sim \sum_{i=1}^n \overline{H}_i(x). \quad (7)$$

Corollary 1. (1) Under the conditions of Theorem [1](#), if $F \in R_{-\alpha}$ for some $\alpha \geq 0$, then, for each fixed $n \geq 1$, it holds that

$$\psi(x, n) \sim \mathbf{E}Y_c^\alpha \frac{1 - (\mathbf{E}Y^\alpha)^n}{1 - \mathbf{E}Y^\alpha} \overline{F}(x), \quad (8)$$

by convention, $(1 - (\mathbf{E}Y^\alpha)^n)/(1 - \mathbf{E}Y^\alpha) = n$ if $\mathbf{E}Y^\alpha = 1$.

Theorem 2. *Under the conditions of Theorem 1, if $J_F^- > 0$ and $\mathbf{E}Y^p < 1$ for some $p > J_F^+$, then it holds that*

$$\psi(x) \sim \sum_{i=1}^{\infty} \overline{H}_i(x). \quad (9)$$

Corollary 2. *Under the conditions of Theorem 2, if $F \in R_{-\alpha}$ for some $\alpha > 0$, then*

$$\psi(x) \sim \frac{\mathbf{E}Y_c^\alpha}{1 - \mathbf{E}Y^\alpha} \overline{F}(x). \quad (10)$$

Moreover, (8) holds uniformly over the integers $\{n \geq 1\}$.

The last result shows that the asymptotic relation (11) for the finite time probability is uniform over the integers $\{n \geq 1\}$.

Theorem 3. *Under the conditions of Theorem 2, the asymptotic relation*

$$\psi(x, n) \sim \sum_{i=1}^n \overline{H}_i(x) . \quad (11)$$

holds uniformly over the integers $\{n \geq 1\}$.

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Thank you!