

# Levy Subordinated Hierarchical Archimedean Copula : Theory and Application

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- Elliptical dependence structures (Gaussian and Student's t Copulas) are widely used.  
 $\Rightarrow$  do not allow for correlation asymmetries
- Archimedean Copulas (AC) such as Gumbel and Calyton Copula are used to capture the asymmetric tail dependence.

# Archimedean copulas

$$C(u_1, u_2, \dots, u_d) = \psi[\psi^{-1}(u_1) + \dots + \psi^{-1}(u_d)],$$

is a  $d$ -dimensional Archimedean copula iff  $\psi \in \mathcal{G}$  defined as

$$\left\{ \psi : [0, \infty) \rightarrow [0, 1] \mid \psi(0) = 1, \psi(\infty) = 0, (-1)^k \frac{d^k}{du^k} \psi(u) \geq 0, k \in \mathbb{N} \right\}$$

a set of *completely monotonic (c.m.)* functions.

# Drawbacks of AC

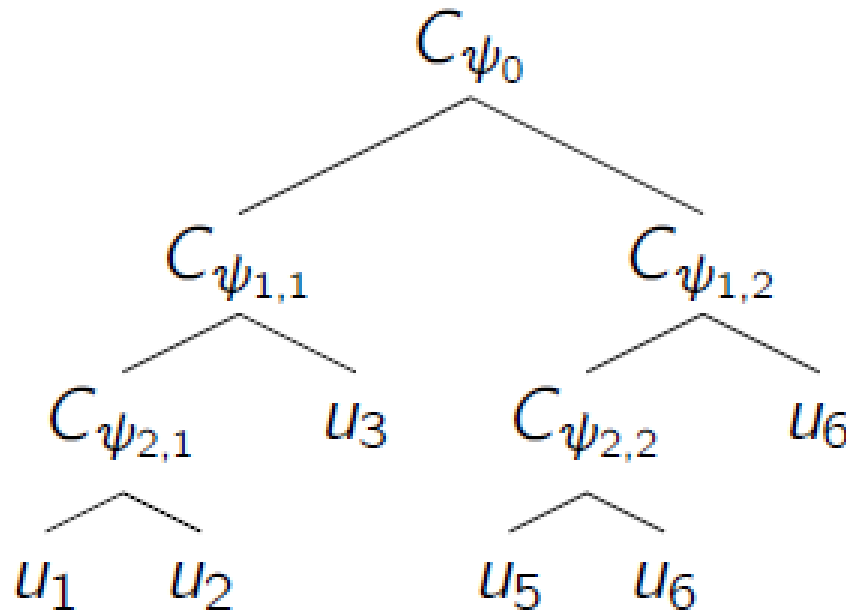
## Exchangeable

- invariant under permutation

=> inappropriate in the multivariate model.

- Hierarchical Archimedean copula (HAC) has been proposed to overcome the disadvantage of Archimedean Copulas (AC).

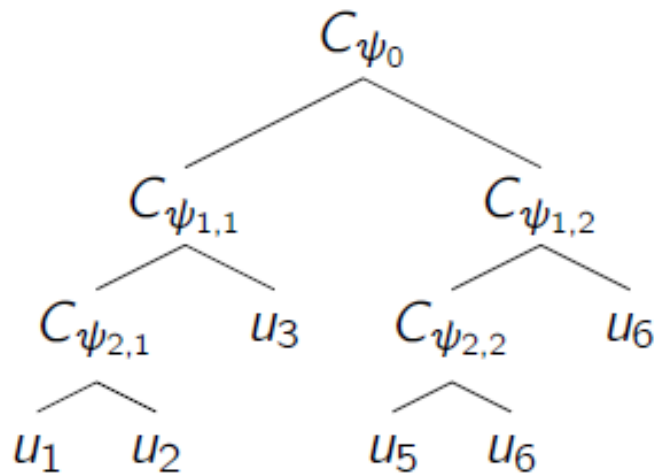
# 3-level, six-dimensional HAC



## “Copula of Copulas” Structure

$$C_{0,1}^{(0)}(u_1, \dots, u_6)$$

$$= C_{0,1}^{(0)}(C_{1,1}^{(1)}(C_{1,1}^{(2)}(u_1, u_2), u_3), C_{1,2}^{(1)}(C_{2,1}^{(2)}(u_4, u_5), u_6))$$



$$\psi_0, \psi_{1,1}, \psi_{1,2}, \psi_{2,1}, \psi_{2,2} \in \mathcal{G}, \quad (3)$$

$$(\psi_0^{-1} \circ \psi_{1,j})' \text{ and } (\psi_{1,i}^{-1} \circ \psi_{2,j})' \in \mathcal{G}, \quad (4)$$

$$i = 1, 2, \quad j = 1, 2 \quad (5)$$

- Condition (4) is called **compatible condition**, adding difficulties in constructing HAC.
- Gumbel HAC satisfies Equation (4) and are easy to simulate. As a result, it is the most frequently used HAC in empirical study.

# Levy Subordinated HAC

- Hering et al. (2010) first construct a two-level hierarchical model based on Lévy subordinators.

Theorem 2.1:

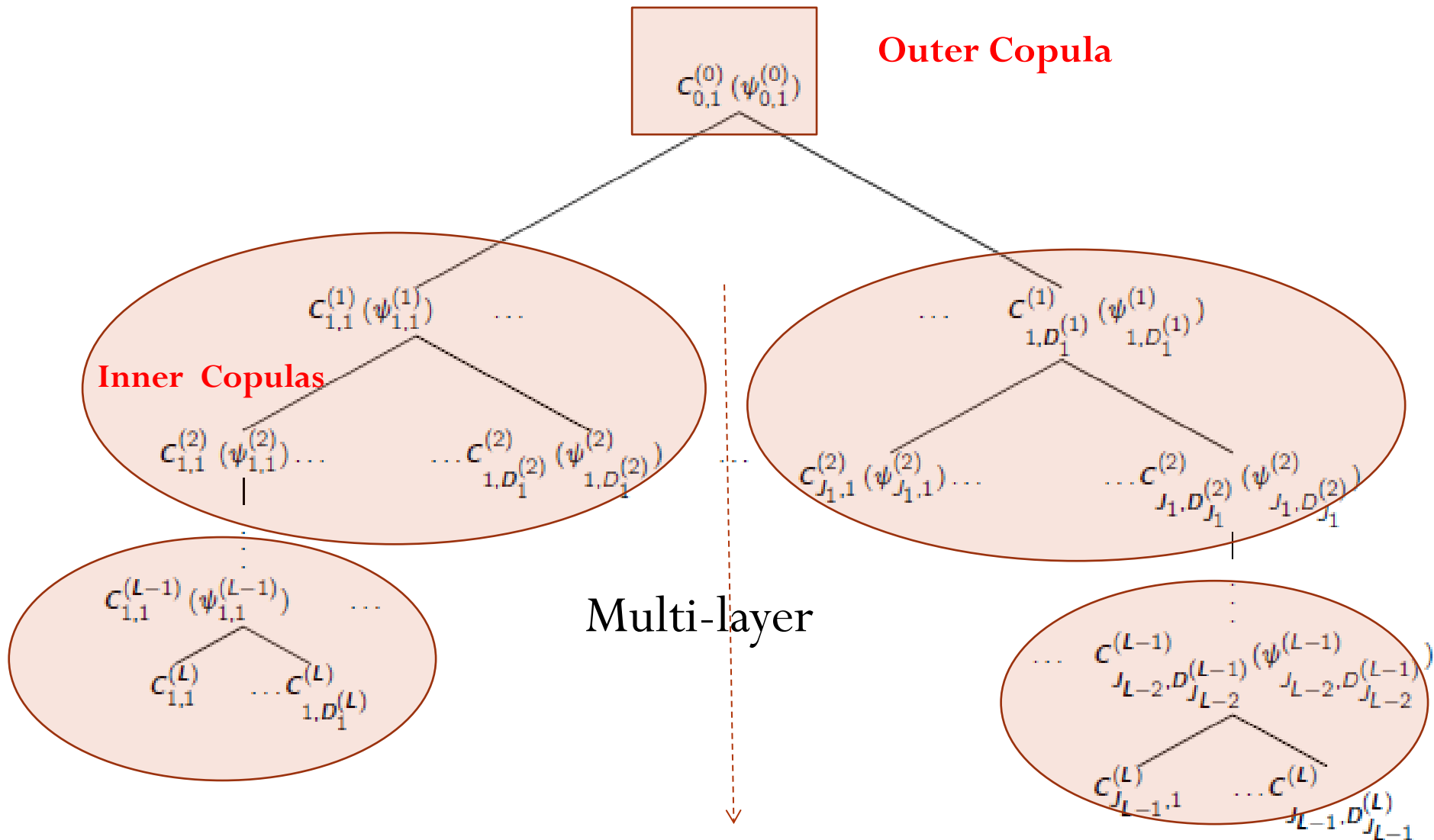
$(\psi_0^{-1} \circ \psi_1)'$  is c.m.  $\Leftrightarrow \psi_0^{-1} \circ \psi_1 = \Psi$  is the Laplace exponent of a Lévy subordinator.

- Mai and Scherer (2012), introduced h-extendible copulas including a three-level LS-HAC case.

# Main Goals

- Introduce a general framework of LS-HAC models with **arbitrary levels**.
- Propose a **three-stage** estimation procedure. We use hierarchical clustering analysis to determine LSHAC structure.
- **Empirically** examine performance of LS-HAC models.

Figure : General Framework of LS-HAC Model



# Integral Representation $\Rightarrow$ For Simulation Purpose

## Theorem

Given the general structure of LS-HAC, the copula function can be constructed as the following:

$$\int_0^\infty \prod_{j_1=1}^{D_{s_0}^{(1)}} \int_0^\infty \prod_{j_2=1}^{D_{s_1}^{(2)}} \int_0^\infty \cdots \int_0^\infty \prod_{j_{L-1}=1}^{D_{s_{L-2}}^{(L-1)}} \int_0^\infty \prod_{j_L=1}^{D_{s_{L-1}}^{(L)}} \left( F_{s_{L-2}j_{L-1}}^{(L-1)}(\bar{u}_{s_{L-1}j_L}) \right)^{v_{s_{L-2}j_{L-1}}^{(L-1)}} (dG)_{j_{L-1}}^{(L-1)},$$

where we define ( $l = 1, \dots, L-1$ ):

$$(dG)_{j_0}^{(0)} = dG_{0,1}^{(0)}(v_{0,1}^{(0)}),$$

and

$$(dG)_{j_l}^{(l)} = d\tilde{G}_{s_{l-1}j_l}^{(l)}(v_{s_{l-1}j_l}^{(l)}; v_{s_{l-2}j_{l-1}}^{(l-1)}) \cdots d\tilde{G}_{s_0j_1}^{(l)}(v_{s_0j_1}^{(1)}; v_{0,1}^{(0)}) dG_{0,1}^{(0)}(v_{0,1}^{(0)}).$$

# Generators of LS-HAC

Proposition  $\Rightarrow$  **For Calibration Purpose**

For  $1 \leq l \leq L$ , in level  $l$ , the  $j_l$ -th copula generator in position  $s_{l-1}$ :  $\psi_{s_{l-1}j_l}^{(l)}$ , can be expressed as:

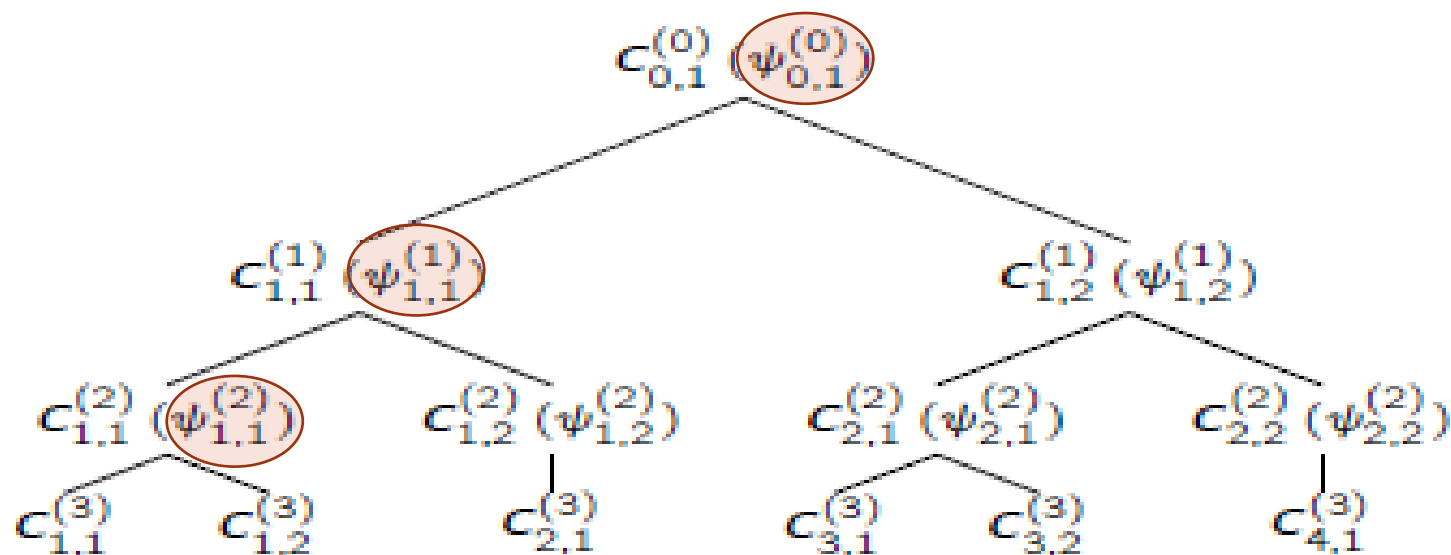
$$\psi_{s_{l-1}j_l}^{(l)} = \psi_{0,1}^{(0)} \bigodot_{i=1}^{l-1} \tilde{\psi}_{s_i j_{i+1}}^{(i+1)} \quad (6)$$

**Outer generator**
**Lévy subordinators**

**Inner generator**

where  $\bigodot_{i=1}^n f_i := f_1 \circ \dots \circ f_n$ , and  $\psi_{s_{l-1}j_l}^{(l)}$  is c.m..

# A three-layer Six-Dimensional LS-HAC



$$\psi_{1,1}^{(1)} = \psi_{0,1}^{(0)} \circ \tilde{\Psi}_{1,1}^{(1)},$$

$$\psi_{1,1}^{(2)} = \psi_{1,1}^{(1)} \circ \tilde{\Psi}_{1,1}^{(2)} = \psi_{0,1}^{(0)} \circ \tilde{\Psi}_{1,1}^{(1)} \circ \tilde{\Psi}_{1,1}^{(2)}$$

In the previous expressions,  $\psi_{i,j}^{(l)}$  is AC generators and  $\tilde{\Psi}_{s,t}^{(l)}$  is Laplace exponent of Lévy subordinators.

# LS-HAC

In this paper, we consider

- AC generators
- Clayton (CL)
- Gumbel (GM)
- Inverse Gaussian (IG)

Name	$\psi(u)$
CL	$(1 + u)^{-\frac{1}{\theta}}, \theta \geq 0$
GM	$\exp(-x^{\frac{1}{\theta}})$
IG	$\exp(\frac{1}{\theta}(1 - \sqrt{1 + 2\theta^2 x}))$

# LS-HAC

- Levy subordinators
  - Gamma (G)
  - Gumbel or Stable (GM)
  - Inverse Gaussian (IG)

Distribution	Inverse Gaussian (IG)	Gamma (G)
Characteristic Function (c.f.)	$\phi_{IG}(u) = \exp\left(-\omega(a\sqrt{2ui + b^2} - ab)\right)$	$\phi_G = \exp\left(-\omega a \log(1 - iu/b)\right)$
Laplace Exponent	$\Psi_{IG}(u) = a\sqrt{2u + b^2} - ab$	$\Psi_G(u) = a \log(1 + u/b)$

- *Note that Gumbel HAC is a special case of LS-HAC when outer generator and subordinator are Gumbel.*

# LS-HAC Estimation

- **Three-Stage Estimation**
  - Stage 1: Estimating Margins
  - **Stage 2: Determining Hierarchical Structure  
=> Hierarchical Clustering Analysis**
  - Stage 3: Estimating LS-HAC parameters by using global MLE, instead of using sequential MLE in classical Gumbel HAC.

# Stage 1: ARMA-GARCH-GH models

- Autoregressive moving average-generalized autoregressive conditional heteroskedasticity (ARMA-GARCH) models (Bollerslev, 1986; Engle, 1982) are used to analyze the dynamics of the margins for the financial data.

## ARMA (m,n)-GARCH (p,q) Model

$$r_{i,t} = c_i + \sum_{j=1}^m \phi_{i,j} r_{i,t-j} + \sum_{j=1}^n \theta_{i,j} \epsilon_{i,t-j} + \epsilon_{i,t},$$

$$\epsilon_{i,t} = \sqrt{h_{i,t}} z_{i,t},$$

$$h_{i,t} = \omega_i + \sum_{j=1}^p a_{i,j} h_{i,t-j} + \sum_{j=1}^q b_{i,j} \epsilon_{i,t-j}^2.$$

# Generalized Hyperbolic (GH)

- Barndorff-Nielsen (1977) uses the GH distribution for financial time series exhibiting skewness, leptokurtosis and tail-thickness.
- The GH distribution nests many well-known highly flexible distributions (normal, Student- $t$ , hyperbolic, Variance Gamma (VG), Normal Inverse Gaussian and GH skewed  $t$  distributions as special case or limit cases.
- After obtaining the marginal parameters, we use probability transform based on the best goodness-of-fit residual distribution to obtain the inputs of copulas.

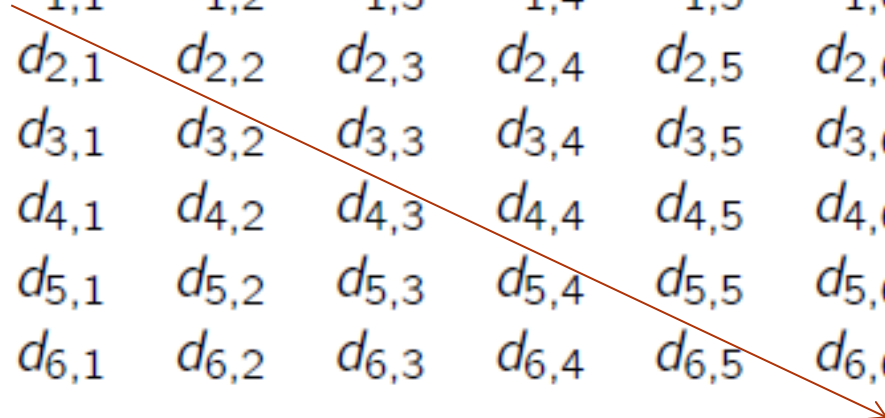
$$u_{it} = F_{Z_i}(z_{it})$$

## Stage 2: Hierarchical Clustering Analysis

- Despite its importance, few papers discuss the grouping method of the HAC or LSHAC models.
- Okhrin et. al. (2013) introduce an estimation procedure to determine optimal structure of HAC by evaluating all possible structures.
  - Time-consuming and inefficient to enumerate all possible structures, especially for high-dimensional problems.
- We employ hierarchical clustering method to determine the structure of LSHAC.

# Dissimilarity (Proximity) Matrix

Using six dimension data as an example:

$$\begin{matrix} & x_1 & x_2 & x_3 & x_4 & x_5 & x_6 \\ \begin{pmatrix} d_{1,1} & d_{1,2} & d_{1,3} & d_{1,4} & d_{1,5} & d_{1,6} \\ d_{2,1} & d_{2,2} & d_{2,3} & d_{2,4} & d_{2,5} & d_{2,6} \\ d_{3,1} & d_{3,2} & d_{3,3} & d_{3,4} & d_{3,5} & d_{3,6} \\ d_{4,1} & d_{4,2} & d_{4,3} & d_{4,4} & d_{4,5} & d_{4,6} \\ d_{5,1} & d_{5,2} & d_{5,3} & d_{5,4} & d_{5,5} & d_{5,6} \\ d_{6,1} & d_{6,2} & d_{6,3} & d_{6,4} & d_{6,5} & d_{6,6} \end{pmatrix} & x_1 & x_2 & x_3 & x_4 & x_5 & x_6 \end{matrix}$$


- $d_{i,j}$  represents dissimilarity of  $i$  and  $j$ .
- $d_{i,i}=0$  and  $d_{i,j} = d_{j,i}$  (symmetric)
- High  $d_{i,j}$  represents high degree of dissimilarity

# Euclidean V.S. Dependence Distance:


- Euclidean Distance:  $d_{ij}^{\text{Euclid}} = \sqrt{\sum_k (x_{k,i} - x_{k,j})^2}$
- Dependence Distance  $d_{ij}^{\text{Corr}} = 1 - \mathfrak{R}_{ij}$

where  $\mathfrak{R}_{ij}$  is an **association measure**, such as *Spearman's rho* and *Kendall's tau*.

- In literature, a sequential procedure according to Kendall's tau is used to determine hierarchical structure of classical Gumbel HAC.
- It ignores *dissimilarity* from “Euclidean Distance”.

# Counter Example

- This example shows the problem of the two distances.

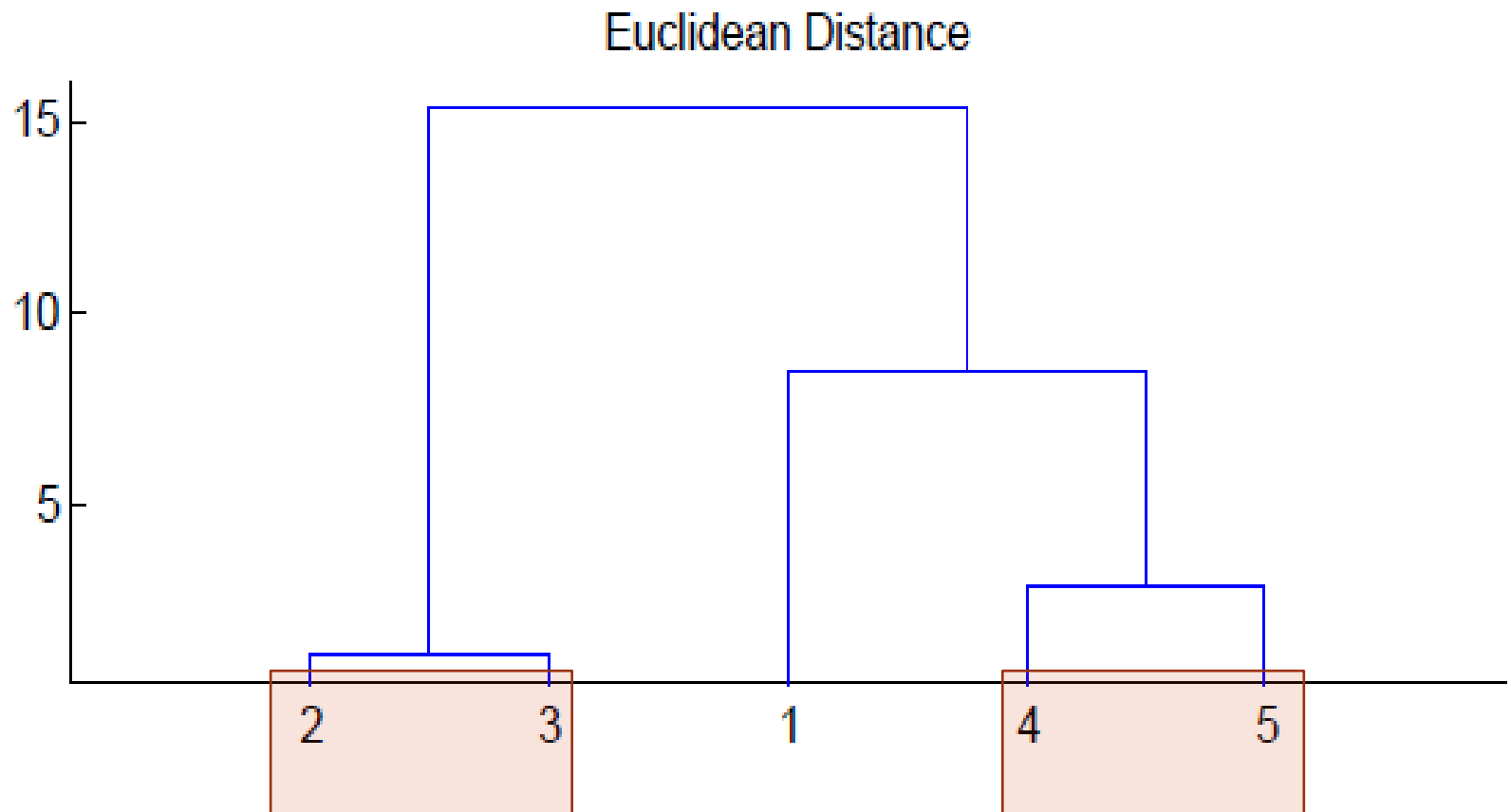


A red bracket on the left side of the equations indicates that  $x_1$  is mapped to the group containing  $x_2, x_3, x_4,$  and  $x_5$ .

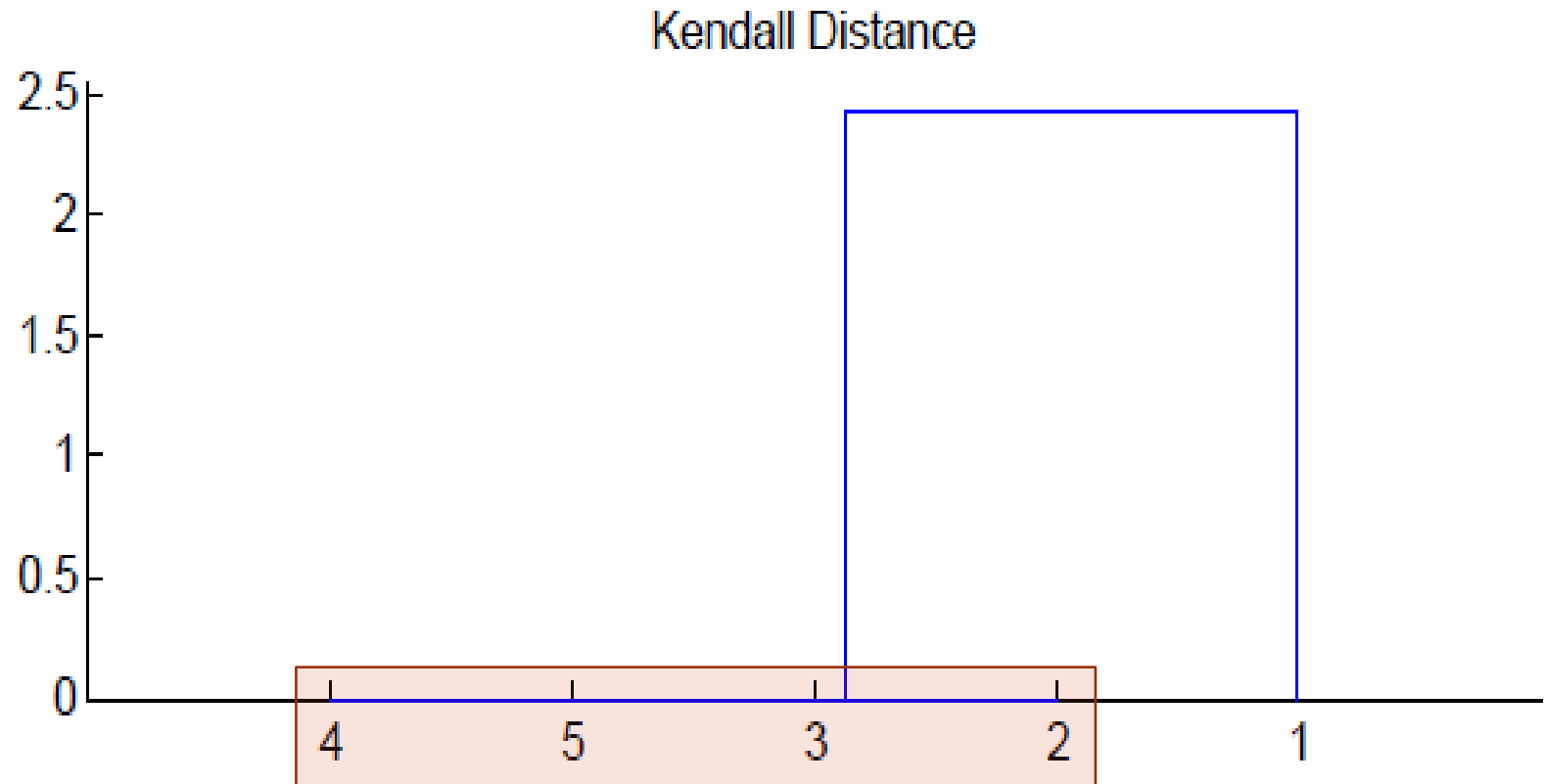
$$\begin{aligned}x_1 &= \text{rand}(1, 100), \\x_2 &= 0.051 : 0.001 : 0.15, \\x_3 &= 0.001 : 0.001 : 0.10, \\x_4 &= 0.851 : 0.001 : 0.95, \\x_5 &= 0.9 * (0.801 : 0.001 : 0.90)\end{aligned}$$

- Intuitively, we have
  - $x_2$  and  $x_3 \Rightarrow$  perfect positive correlated with small values
  - $x_4$  and  $x_5 \Rightarrow$  perfect positive correlated with large values
  - $x_1 \Rightarrow$  independent of others (third group)

# Euclidean Distance



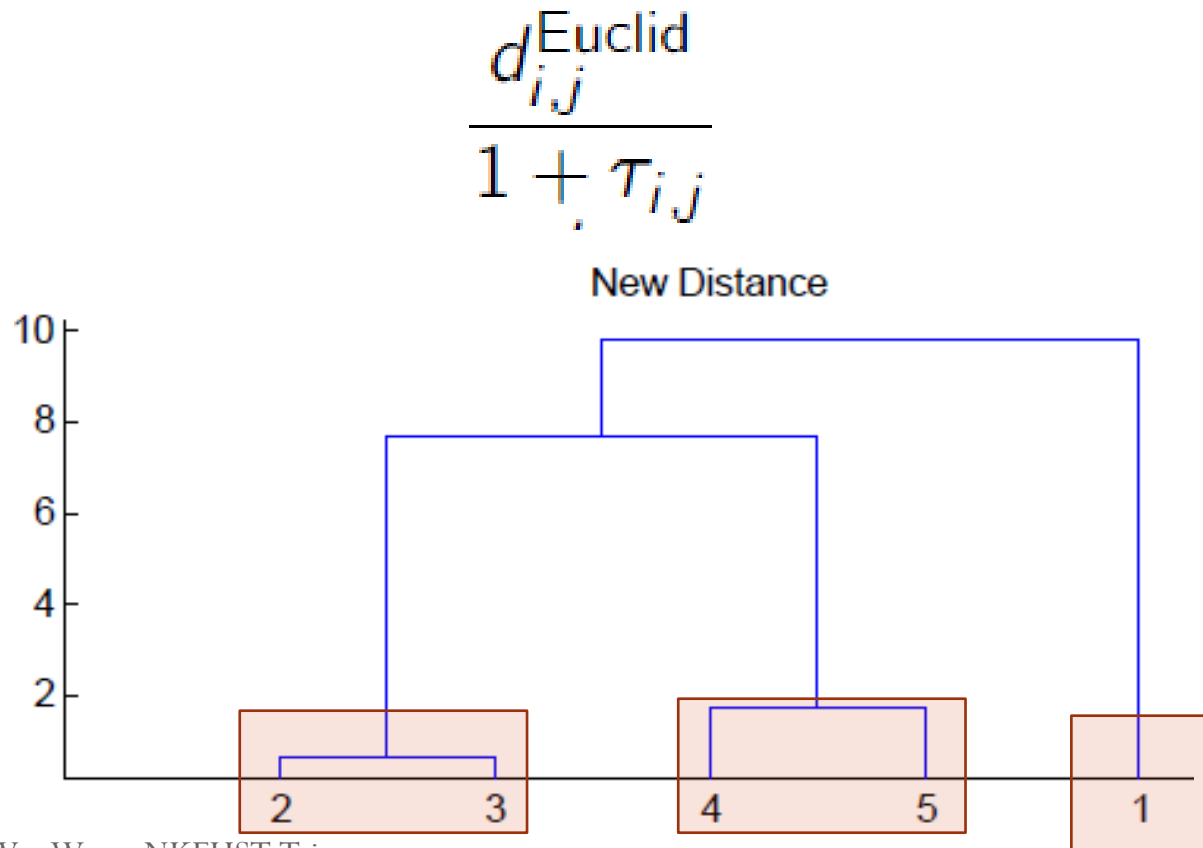
# Kendall's Tau measure



It groups  $x_2$  to  $x_5$  as a *perfect-positive-correlated* group without further discrimination

# A Trade-off measure

- In this paper, we propose a new measure by dividing Euclidean distance by dependence distances



# Empirical Analysis

- Data: 5-year daily data of Stock indices
  - European (France, UK and Germany)
  - American (Brazil, Canada and US)

# Stage 1: Margins

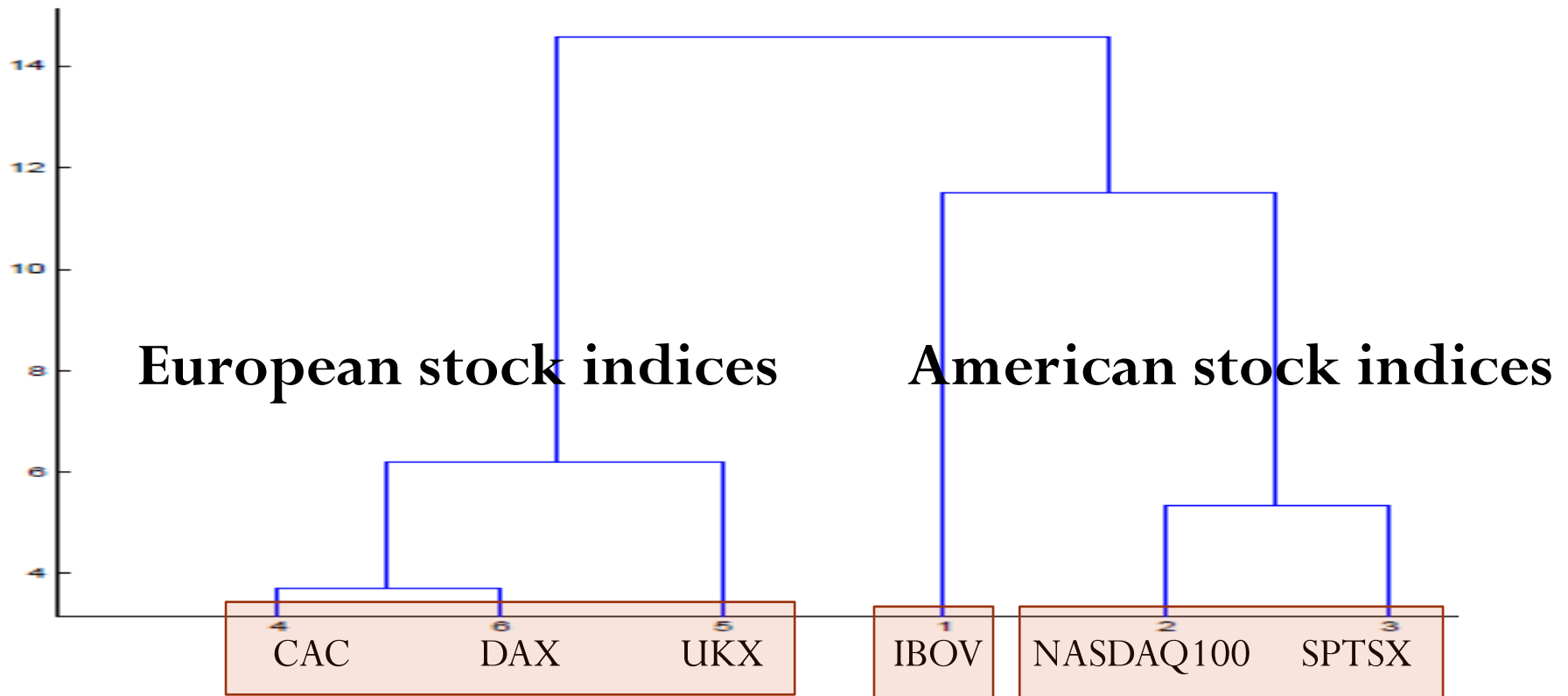
Marginal Estimation Results (GARCH(1,1) effect but no ARMA effect)

Indices	IBOV	NDX	SPTSX	CAC	UKX	DAX
Best Distribution	t	VG	VG	t	VG	VG
$c(\times 10^{-4})$	1.7676 (0.3753)	10.4844 (3.8384)***	2.4634 (0.8901)	9.5697 (2.4141)***	4.5534 (1.5001)	8.8959 (2.2612)**
$\omega(\times 10^{-6})$	5.617 (2.227)**	2.9194 (2.7788)***	3.9638 (0.9451)	2.7935 (1.526)*	8.2843 (1.2079)	1.3621 (1.1318)
$a_1$	0.9191 (53.6884)***	0.8775 (36.0184)***	0.9302 (62.5136)***	0.9254 (46.6514)***	0.9374 (51.7158)***	0.9363 (58.6861)***
$b_1$	0.0685 (4.3703)***	0.1048 (4.5191)***	0.0698 (4.3598)***	0.0683 (3.5849)***	0.0586 (3.3877)***	0.0597 (3.9366)***
$\nu$	8 (6.7306)***			6.5103 (3.5849)***		
$\alpha$		1.9846 (11.8214)***	2.4289 (7.3098)***		2.4826 (7.5641)***	2.174 (9.3175)***
$\beta$		-0.1753 (11.8214)***	-0.415 (-3.3496)***		-0.1739 (-1.9242)***	-0.1739 (-1.9242)***

# Empirical Analysis: Hierarchical clustering

Dendrogram according to our distance measure

## Geographic Groups



# Empirical Analysis: Global MLE

Table : Six Dimension Estimating Results

	Copula	Log-Likelihood	AIC	BIC
Elliptical	Gaussian	3361.4681	-3346.4681	-3309.3370
	TV-Gaussian	3384.0908	-3367.0908	-3325.0089
	T	3558.7891	-3542.7891	-3503.1826
	TV-T	3573.9428	-3555.9428	-3511.3855
LS-HAC	GM-family	3566.5219	-3557.5219	-3534.2425
	CL-family	3622.3425	-3613.1	-3590.1
	IG-family	3631.47197	-3622.472	-3599.1926
	GM-GM-GM	3298.2271	-3293.2271	-3280.2941

Note: "TV-" represents time-varying copula models. GM represents Gumbel copula, CL represents Clayton copula, IG represents IG copula, G represents Gamma Lévy exponent, and IG represents Inverse Gaussian Lévy exponent.

# Conclusions

- Verify the multi-layer LSHAC models theoretically.
- Introduce a three-stage estimation procedure and provide a new dissimilarity measure by incorporating Euclidean and dependence distances.
- Demonstrate that compared to time-varying elliptical copulas, LS-HAC models provide the best goodness of fit.

# Thanks for Attentions!

# Six-Dimensional LS-HAC

- $\psi_{0,1}^{(0)}$  (outer generator) : GM or  $GM \circ G$
- If  $\psi_{0,1}^{(0)}$  is  $GM \circ G$  generator and  $\tilde{\psi}_{1,1}^{(1)}$  is IG subordinator, we have a five-parameter inner generator

$$\begin{aligned}\psi_{1,1}^{(1)} &= GM \circ G \circ IG \\ &= \exp\left(-\left(a_1 \log\left(1 + \frac{a_2}{b_1}(\sqrt{2u + b_2^2} - b_2)\right)\right)^{\frac{1}{\theta}}\right)\end{aligned}$$

- LS-HAC model is very flexible.
- *Note that Gumbel HAC is a special case of LS-HAC when outer generator and subordinator are Gumbel.*

# References

- [1] Lee, S. and Lin, S. (2010), Modeling and Evaluating Insurance Losses via Mixtures of Erlang Distributions. *North American Actuarial Journal* 14(1), 107-130.
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- [4] Christian Hering and Marius Hofert and Jan-Frederik Mai and Matthias Scherer (2010), Constructing Hierarchical Archimedean Copulas with Lévy Subordinators. *Journal of Multivariate Analysis* 101, 1428-1433.
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# Generalized Hyperbolic (GH)

- The probability density function of the GH distribution can be shown as follows

$$f_{GH}(x|\alpha, \beta, \lambda, \delta, m) = \frac{\left(\frac{\sqrt{\alpha^2 - \beta^2}}{\delta}\right)^\lambda}{\sqrt{2\pi} \left(K_\lambda(\delta\sqrt{\alpha^2 - \beta^2})\right)} e^{\beta(z_t - m)} \frac{K_{\lambda - \frac{1}{2}}\left(\alpha\sqrt{\delta^2 + (x - m)^2}\right)}{\left(\frac{\sqrt{\delta^2 + (x - m)^2}}{\alpha}\right)^{\frac{1}{2} - \lambda}},$$

- The moment generating function of GH distribution is given by

$$\phi_{GH}(\omega) = \exp(\psi_{GH}(\omega)) = e^{i\omega m} \left(-\frac{\alpha^2 - \beta^2}{\alpha^2 - (\beta + i\omega)^2}\right)^{\lambda/2} \frac{K_\lambda(\delta\sqrt{\alpha^2 - (\beta + i\omega)^2})}{K_\lambda(\delta\sqrt{\alpha^2 - \beta^2})}$$